Regret Analysis of Stochastic and Nonstochastic Multi-armed Bandit Problems, Part I

> Sébastien Bubeck Theory Group

Research



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Protocol: For each round t = 1, 2, ..., T, the player chooses $I_t \in [n]$ based on past observations and receives a reward/observation $Y_t \sim \nu_{I_t}$ (independently from the past).

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Performance measure: The cumulative regret is the difference between the player's accumulated reward and the maximum the player could have obtained had she known all the parameters,

$$\overline{R}_T = T\mu^* - \mathbb{E}\sum_{t\in[T]} Y_t.$$

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Fundamental tension between **exploration** and **exploitation**. Many applications!

How small can we expect \overline{R}_T to be? Consider the 2-armed case where $\nu_1 = \text{Ber}(1/2)$ and $\nu_2 = Ber(1/2 + \xi\Delta)$ where $\xi \in \{-1, 1\}$ is unknown.

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With τ expected observations from the second arm there is a probability at least $\exp(-\tau\Delta^2)$ to make the wrong guess on the value of ξ .

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$$\overline{R}_{T}(\xi = +1) + \overline{R}_{T}(\xi = -1) \geq \Delta \tau(T) + \Delta \sum_{t=1}^{r} \exp(-\tau(t)\Delta^{2})$$
$$\geq \Delta \min_{t \in [T]} (t + T \exp(-t\Delta^{2}))$$
$$\approx \frac{\log(T\Delta^{2})}{\Delta}.$$

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See Bubeck, Perchet and Rigollet [2012] for the details.

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See Bubeck, Perchet and Rigollet [2012] for the details. For Δ fixed the lower bound is $\frac{\log(T)}{\Delta}$, and for the worse Δ ($\approx 1/\sqrt{T}$) it is \sqrt{T} (Auer, Cesa-Bianchi, Freund and Schapire [1995]: \sqrt{Tn} for the *n*-armed case).

Notation: $\Delta_i = \mu^* - \mu_i$ and $N_i(t)$ is the number of pulls of arm *i* up to time *t*. Then one has $\overline{R}_T = \sum_{i=1}^n \Delta_i \mathbb{E} N_i(T)$.

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, $kl(p, q) := p \log \frac{p}{q} + (1 - p) \log \frac{1 - p}{1 - q}$.

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Theorem (Lai and Robbins [1985]) Consider a strategy s.t. $\forall a > 0$, we have $\mathbb{E}N_i(T) = o(T^a)$ if $\Delta_i > 0$. Then for any Bernoulli distributions,

$$\liminf_{n\to+\infty} \frac{\overline{R}_T}{\log(T)} \geq \sum_{i:\Delta_i>0} \frac{\Delta_i}{\operatorname{kl}(\mu_i,\mu^*)}.$$

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Note that $\frac{1}{2\Delta_i} \ge \frac{\Delta_i}{\mathrm{kl}(\mu_i,\mu^*)} \ge \frac{\mu^*(1-\mu^*)}{2\Delta_i}$ so up to a variance-like term the Lai and Robbins lower bound is $\sum_{i:\Delta_i>0} \frac{\log(T)}{2\Delta_i}$.

i.i.d. multi-armed bandit: fundamental strategy Hoeffding's inequality: w.p. $\geq 1 - 1/T$, $\forall t \in [T], i \in [n]$,

$$\mu_i \leq \frac{1}{N_i(t)} \sum_{s < t: I_s = i} Y_s + \sqrt{\frac{2\log(T)}{N_i(t)}} =: \mathrm{UCB}_i(t).$$

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$$I_t \in \underset{i \in [n]}{\operatorname{argmax}} \operatorname{UCB}_i(t).$$

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so that $\mathbb{E}N_i(T) \leq 2 + 8\log(T)/\Delta_i^2$ and in fact

$$\overline{R}_{\mathcal{T}} \leq 2 + \sum_{i:\Delta_i > 0} \frac{8\log(\mathcal{T})}{\Delta_i}.$$

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i.i.d. multi-armed bandit: going further

 Optimal constant (replacing 8 by 1/2 in the UCB regret bound) and Lai and Robbins variance-like term (replacing Δ_i by kl(μ_i, μ^{*})): see Cappé, Garivier, Maillard, Munos and Stoltz [2013].

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- 2. In many applications one is merely interested in *finding* the best arm (instead of maximizing cumulative reward): this is the best arm identification problem. For the fundamental strategies see Even-Dar, Mannor and Mansour [2006] for the fixed-confidence setting (see also Jamieson and Nowak [2014] for a recent short survey) and Audibert, Bubeck and Munos [2010] for the fixed budget setting. Key takeaway: one needs of order $\mathbf{H} := \sum_i \Delta_i^{-2}$ rounds to find the best arm.

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- The UCB analysis extends to sub-Gaussian reward distributions. For heavy-tailed distributions, say with 1 + ε moment for some ε ∈ (0,1], one can get a regret that scales with Δ_i^{-1/ε} (instead of Δ_i⁻¹) by using a robust mean estimator, see Bubeck, Cesa-Bianchi and Lugosi [2012].

Adversarial multi-armed bandit, Auer, Cesa-Bianchi, Freund and Schapire [1995, 2001]

For t = 1, ..., T, the player chooses $I_t \in [n]$ based on previous observations, and simultaneously an adversary chooses a loss vector $\ell_t \in [0, 1]^n$. The player's loss/observation is $\ell_t(I_t)$.

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$$R_{\mathcal{T}} = \max_{i \in [n]} \sum_{t \in [\mathcal{T}]} (\ell_t(I_t) - \ell_t(i)), \quad \overline{R}_{\mathcal{T}} = \max_{i \in [n]} \mathbb{E} \sum_{t \in [\mathcal{T}]} (\ell_t(I_t) - \ell_t(i)).$$

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Obviously $\mathbb{E}R_T \geq \overline{R}_T$ and there is equality in the oblivious case (\equiv adversary's choice are independent of the player's choice). The case where ℓ_1, \ldots, ℓ_T is an i.i.d. sequence corresponds to the i.i.d. case we just studied. In particular we have a \sqrt{Tn} lower bound.

Exponential weights strategy for *full information* (ℓ_t is observed at the end of round *t*): play I_t at random from p_t where

$$p_{t+1}(i) = \frac{1}{Z_{t+1}} p_t(i) \exp(-\eta \ell_t(i)).$$

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$$\eta \sum_{t} \left(\sum_{i} p_{t}(i) \ell_{t}(i) - \ell_{t}(j) \right) = \operatorname{Ent}(\delta_{j} \| p_{1}) - \operatorname{Ent}(\delta_{j} \| p_{T+1}) + \sum_{t} \psi_{t}(j)$$

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$$\eta \sum_{t} \left(\sum_{i} p_{t}(i) \ell_{t}(i) - \ell_{t}(j) \right) = \operatorname{Ent}(\delta_{j} \| p_{1}) - \operatorname{Ent}(\delta_{j} \| p_{T+1}) + \sum_{t} \psi_{t}$$
$$\operatorname{Using that} \ell_{t} \geq 0 \text{ one has } \psi_{t} \leq \frac{\eta^{2}}{2} \mathbb{E} \ell_{t}(i)^{2} \text{ thus } \overline{R}_{T} \leq \frac{\log(n)}{\eta} + \frac{\eta T}{2}$$

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Exp3: replace ℓ_t by $\widetilde{\ell}_t$ in the exponential weights strategy, where

$$\widetilde{\ell}_t(i) = \frac{\ell_t(I_t)}{p_t(i)} \mathbb{1}\{i = I_t\}.$$

Key property: $\mathbb{E}_{I_t \sim p_t} \tilde{\ell}_t(i) = \ell_t(i).$



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$$\overline{R}_{\mathcal{T}} \leq \frac{\log(n)}{\eta} + \frac{\eta}{2} \mathbb{E} \sum_{t} \mathbb{E}_{I \sim p_{t}} \widetilde{\ell}_{t}(I)^{2}.$$

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Amazingly the variance term is automatically controlled:

$$\mathbb{E}_{I_t,I\sim p_t}\widetilde{\ell}_t(I)^2 \leq \mathbb{E}_{I_t,I\sim p_t}\frac{\mathbb{1}\{I=I_t\}}{p_t(I_t)^2} = \mathbb{E}_{I\sim p_t}\frac{1}{p_t(I)} = n.$$

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Thus with $\eta = \sqrt{2n \log(n)/T}$ one gets $\overline{R}_T \leq \sqrt{2Tn \log(n)}$.

1. With the modified loss estimate $\frac{\ell_t(l_t)\mathbb{1}\{i=l_t\}+\beta}{p_t(l_t)}$ one can prove high probability bounds on R_T , and by integrating the deviations one can show $\mathbb{E}R_T = O(\sqrt{Tn\log(n)})$.

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4. There exists strategies which guarantee simultaneously $\overline{R}_T = \widetilde{O}(\sqrt{Tn})$ in the adversarial model and $\overline{R}_T = \widetilde{O}(\sum_i \Delta_i^{-1})$ in the i.i.d. model, see Bubeck and Slivkins [2012].

- 1. With the modified loss estimate $\frac{\ell_t(l_t)\mathbb{1}\{i=l_t\}+\beta}{p_t(l_t)}$ one can prove high probability bounds on R_T , and by integrating the deviations one can show $\mathbb{E}R_T = O(\sqrt{Tn\log(n)})$.
- 2. The extraneous logarithmic factor in the pseudo-regret upper can be removed, see Audibert and Bubeck [2009]. Conjecture: one cannot remove the log factor for the expected regret, that is for any strategy there exists an adaptive adversary such that $\mathbb{E}R_T = \Omega(\sqrt{Tn \log(n)}).$
- 3. *T* can be replaced by various measure of "variance" in the loss sequence, see e.g., Hazan and Kale [2009].
- 4. There exists strategies which guarantee simultaneously $\overline{R}_T = \widetilde{O}(\sqrt{Tn})$ in the adversarial model and $\overline{R}_T = \widetilde{O}(\sum_i \Delta_i^{-1})$ in the i.i.d. model, see Bubeck and Slivkins [2012].
- 5. Graph feedback structure, regret with respect to S switches, label efficient, switching cost...

Set of models $\{(\nu_1(\theta), \ldots, \nu_n(\theta)), \theta \in \Theta\}$ and prior distribution π_0 over Θ . The Bayesian regret is defined as

$$BR_T(\pi_0) = \mathbb{E}_{\theta \sim \pi_0} \overline{R}_T(\nu_1(\theta), \dots, \nu_n(\theta)).$$

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Notation: π_t denotes the posterior distribution on θ at time *t*.

Theorem (Gittins [1979])

Consider the product and γ -discounted case: $\Theta = \times_i \Theta_i$, $\nu_i(\theta) := \nu(\theta_i), \ \pi_0 = \otimes_i \pi_0(i)$, and furthermore one is interested in maximizing $\mathbb{E} \sum_{t \ge 0} \gamma^t Y_t$.

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$$\sup\left\{\lambda\in\mathbb{R}:\sup_{\tau}\mathbb{E}\left(\sum_{t<\tau}\gamma^{t}X_{t}+\frac{\gamma^{\tau}}{1-\gamma}\lambda\right)\geq\frac{1}{1-\gamma}\lambda\right\},$$

where the expectation is over (X_t) drawn from $\nu(\theta)$ with $\theta \sim \pi_s(i)$, and the supremum is taken over all stopping times τ .

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where the expectation is over (X_t) drawn from $\nu(\theta)$ with $\theta \sim \pi_s(i)$, and the supremum is taken over all stopping times τ . For much more (implementation for exponential families, interpretation as a multitoken Markov game, ...) see Dumitriu, Tetali and Winkler [2003], Gittins, Glazebrook, Weber [2011], Kaufmann [2014].

Weber [1992] gives an exquisite proof of Gittins theorem. Let

$$\lambda_t(i) := \sup\left\{\lambda \in \mathbb{R} : \sup_{\tau} \mathbb{E} \sum_{t < \tau} \gamma^t(X_t - \lambda) \ge 0\right\}$$

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- 2. Since the prevailing charge is nonincreasing, the discounted sum of prevailing charge is maximized if we always pick the arm with maximum prevailing charge.
- 3. Gittins index does exactly 2. and that in this case 1. is an equality. Q.E.D.

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Theoretical guarantees for this highly practical strategy have long remained elusive. Recently Agrawal and Goyal [2012] and Kaufmann, Korda and Munos [2012] proved that TS with Bernoulli reward distributions and uniform prior on the parameters achieves $\overline{R}_{T} = O\left(\sum_{i} \frac{\log(T)}{\Delta_{i}}\right)$ (note that this is the frequentist regret!).

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Guha and Munagala [2014] conjecture that, for product priors, TS is a 2-approximation to the optimal Bayesian strategy for the objective of minimizing the number of pulls on suboptimal arms.

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Assume a prior in the adversarial model, that is a prior over $(\ell_1, \ldots, \ell_T) \in [0, 1]^{n \times T}$, and let \mathbb{E}_t denote the posterior distribution (given $\ell_1(I_1), \ldots, \ell_{t-1}(I_{t-1})$).

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 $r_t(i) = \mathbb{E}_t(\ell_t(i) - \ell_t(i^*)), \text{ and } v_t(i) = \operatorname{Var}_t(\mathbb{E}_t(\ell_t(i)|i^*)).$

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Key observation (next slide):

$$\mathbb{E}\sum_{t\leq T} v_t(I_t) \leq \frac{1}{2}H(x^*)$$

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which implies:

$$\forall t, \mathbb{E}_t r_t(I_t) \leq \sqrt{C \ \mathbb{E}_t v_t(I_t)}$$

$$\Rightarrow \ \mathbb{E} \sum_{t=1}^T r_t(I_t) \leq \sum_{t=1}^T \sqrt{C \ \mathbb{E} v_t(I_t)}$$

$$\Rightarrow \ BR_T \leq \sqrt{C \ T \ H(i^*)/2}.$$

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Bayesian multi-armed bandit, accumulation of information

$$v_t(i) = \operatorname{Var}_t(\mathbb{E}_t(\ell_t(i)|i^*)), \ \pi_t(j) = \mathbb{P}_t(i^* = j), \ \mathbb{E}\sum_{t \leq T} v_t(I_t) \leq \frac{1}{2}H(x^*)$$

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Equipped with Pinsker's inequality and basic information theory concepts (such as the mutual information \mathbb{I}) one has:

$$\begin{split} v_t(i) &= \sum_j \pi_t(j) (\mathbb{E}_t(\ell_t(i)|i^* = j) - \mathbb{E}_t(\ell_t(i)))^2 \\ &\leq \frac{1}{2} \sum_j \pi_t(j) \mathrm{Ent}(\mathcal{L}_t(\ell_t(i)|i^* = j) \| \mathcal{L}_t(\ell_t(i))) \\ &= \frac{1}{2} \mathbb{I}_t(\ell_t(i), i^*) = H_t(i^*) - H_t(i^*|\ell_t(i)). \end{split}$$

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Thus $\mathbb{E}v_t(I_t) \leq \frac{1}{2}\mathbb{E}(H_t(i^*) - H_{t+1}(i^*)).$

Bayesian multi-armed bandit, TS' information ratio Let $\bar{\ell}_t(i) = \mathbb{E}_t \ell_t(i)$ and $\bar{\ell}_t(i,j) = \mathbb{E}_t(\ell_t(i)|i^* = j)$. Then $\mathbb{E}_t r_t(I_t) \le \sqrt{C \mathbb{E}_t v_t(I_t)}$ $\Leftrightarrow \mathbb{E}_t \bar{\ell}_t(I_t) - \sum_i \pi_t(i) \bar{\ell}_t(i,i) \le \sqrt{C \mathbb{E}_t \sum_j \pi_t(j) (\bar{\ell}_t(I_t,j) - \bar{\ell}_t(I_t))^2}$

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Thus TS always satisfies $BR_T \leq \sqrt{TnH(i^*)} \leq \sqrt{Tn\log(n)}.$

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Thus TS always satisfies $BR_T \leq \sqrt{TnH(i^*)} \leq \sqrt{Tn\log(n)}$. Side

note: by the minimax theorem this implies there exists a strategy for the oblivious adversarial model with regret $\sqrt{Tn\log(n)}$.

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Summary of basic results

- 1. In the i.i.d. model UCB attains a regret of $O\left(\sum_{i} \frac{\log(T)}{\Delta_{i}}\right)$ and by Lai and Robbins' lower bound this is optimal (up to a multiplicative variance term).
- 2. In the adversarial model Exp3 attains a regret of $O(\sqrt{Tn\log(n)})$ and this is optimal up to the logarithmic term.
- 3. In the Bayesian model, Gittins index gives an *optimal* strategy for the case of product priors. For general priors Thompson Sampling is a more flexible strategy. Its Bayesian regret is controlled by the entropy of the optimal decision. Moreover TS with an uninformative prior has frequentist guarantees comparable to UCB.